Multifidelity importance sampling methods for rare event simulation

Benjamin Peherstorfer University of Wisconsin-Madison

Karen Willcox and Boris Kramer MIT

> Max Gunzburger Florida State University

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Introduction

High-fidelity model with costs $w_1 \gg 0$

 $f^{(1)}:\mathcal{D}
ightarrow\mathcal{Y}$

Random variable Z, estimate

 $s = \mathbb{E}[f^{(1)}(Z)]$

Monte Carlo estimator of s with samples Z_1, \ldots, Z_n

$$\bar{y}_n^{(1)} = \frac{1}{n} \sum_{i=1}^n f^{(1)}(Z_i)$$

Computational costs high

- Many evaluations of high-fidelity model
- Typically $10^3 10^6$ evaluations
- Intractable if $f^{(1)}$ expensive



Surrogate models

Given is a high-fidelity model $f^{(1)}: \mathcal{D} \to \mathcal{Y}$

- Large-scale numerical simulation
- Achieves required accuracy
- Computationally expensive

Additionally, often have surrogate models

 $f^{(i)}: \mathcal{D} \to \mathcal{Y}, \qquad i=2,\ldots,k$

- Approximate high-fidelity $f^{(1)}$
- Often orders of magnitudes cheaper

Examples of surrogate models



data-fit models, response surfaces, machine learning



approximations



reduced basis, proper orthogonal decomposition



simplified models, linearized models



Surrogate models in uncertainty quantification

Replace $f^{(1)}$ with a surrogate model

- Costs of uncertainty quantification reduced
- Often orders of magnitude speedups

Estimate depends on surrogate accuracy

- Control with error bounds/estimators
- Rebuild if accuracy too low
- No guarantees without bounds/estimators

Issues

- Propagation of surrogate error on estimate
- Surrogates without error control
- Costs of rebuilding a surrogate model



Our approach: Multifidelity methods

Combine high-fidelity and surrogate models

- Leverage surrogate models for speedup
- Recourse to high-fidelity for accuracy

Multifidelity speeds up computations

- Balance #solves among models
- Adapt, fuse, filter with surrogate models

Multifidelity guarantees high-fidelity accuracy

- Occasional recourse to high-fidelity model
- High-fidelity model is kept in the loop
- Independent of error control for surrogates



[P., Willcox, Gunzburger, Survey of multifidelity methods in uncertainty propagation, inference, and optimization; available online as technical report, MIT, 2016]

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Multifidelity Monte Carlo (MFMC)

MFMC use control variates for var reduction

- Derives control variates from surrogates
- Uses number of samples that minimize error

[P., Willcox, Gunzburger: Optimal model management for multifidelity Monte Carlo estimation. SIAM Journal on Scientific Computing, 2016]

Multifidelity sensitivity analysis

- Identify the parameters of model with largest influence on quantity of interest
- Elizabeth Qian (MIT)/Earth Science (LANL)

Asymptotic analysis of MFMC

[P., Gunzburger, Willcox: Convergence analysis of multifidelity Monte Carlo estimation, submitted. 2016]

MFMC with information reuse

[Ng, Willcox: Monte Carlo Information-Reuse Approach to Aircraft Conceptual Design Optimization Under Uncertainty. 2015]

MFMC with optimally-adapted surrogates



Multifidelity rare event simulation based on importance sampling

with Karen Willcox and Boris Kramer

MFCE: Problem setup

Threshold $0 < t \in \mathbb{R}$ and random variable

 $Z \sim p$

Estimate rare event probability

 $P_f = \mathbb{P}_p[f^{(1)} \le t]$

Can be reformulated to estimating $\ensuremath{\mathbb{E}}$

$$P_f = \mathbb{E}_p[\mathbb{I}_t^{(1)}]$$

with indicator function

$$\mathbb{I}_t^{(1)}(\pmb{z}) = egin{cases} 1, & f^{(1)}(\pmb{z}) \leq t\,, \ 0, & f^{(1)}(\pmb{z}) > t \end{cases}$$

If $P_f \ll 1$, very unlikely that we hit $f^{(1)} \leq t$



MFCE: Importance sampling

Consider a biasing density q with

 $\operatorname{supp}(p) \subseteq \operatorname{supp}(q)$

Reformulate estimation problem

$$P_f = \mathbb{E}_{\rho}[\mathbb{I}_t^{(1)}] = \mathbb{E}_{\boldsymbol{q}}\left[\mathbb{I}_t^{(1)} \frac{p}{q}\right]$$

Goal is constructing suitable q with

$$\mathsf{Var}_{q}\left[\mathbb{I}_{t}^{(1)}\frac{p}{q}\right] \ll \mathsf{Var}_{p}[\mathbb{I}_{t}^{(1)}]$$

 $\Rightarrow \textit{Use surrogate models}$



Two-fidelity approaches

- Switch between models [Li, Xiu et al., 2010, 2011, 2014]
- Reduced basis models with error estimators [Chen and Quarteroni, 2013]
- Kriging models and importance sampling [Dubourg et al., 2013]
- Subset method with machine-learning-based models [Bourinet et al., 2011], [Papadopoulos et al., 2012]
- Surrogates and importance sampling [P., Cui, Marzouk, Willcox, 2016]

Multilevel methods for rare event simulation

- ► Variance reduction via control variates [Elfverson et al., 2016], [Fagerlund et al., 2016]
- Subset method with coarse-grid approximations [Ullmann and Papaioannou, 2015]

Combining multiple general types of surrogates

Importance sampling + control variates [P., Kramer, Willcox, 2017]

MFCE: Direct sampling of surrogate models

Directly sampling surrogate models to construct biasing density

- Reduces costs per sample
- Number of samples to construct biasing density remains the same
- Works well for probabilities $> 10^{-5}$
- \Rightarrow Insufficient for very rare event probabilities in range of 10^{-9}



[P., Cui, Marzouk, Willcox, 2016], [P., Kramer, Willcox, 2017]

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Threshold t controls "rareness" of event

$$\mathbb{I}_t^{(1)}(m{z}) = egin{cases} 1\,, & f^{(1)}(m{z}) \leq t\,, \ 0\,, & f^{(1)}(m{z}) > t \end{cases}$$



[Rubinstein, 1999], [Rubinstein, 2001]

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MFCE: Cross-entropy method – in each step

Need to find biasing density in each step i = 1, ..., T

Optimal biasing density that reduces variance to 0

$$q_i(oldsymbol{z}) \propto \mathbb{I}_{t_i}^{(1)}(oldsymbol{z}) p(oldsymbol{z})$$

 \Rightarrow Unknown normalizing constant (quantity we want to estimate)

▶ Find $q_{v_i} \in \mathcal{Q} = \{q_{v} \ : \ v \in \mathcal{P}\}$ with min Kullback-Leibler distance to q_i

 $\min_{\boldsymbol{v}_i \in \mathcal{P}} D_{\mathsf{KL}}(\boldsymbol{q}_i || \boldsymbol{q}_{\boldsymbol{v}_i})$

▶ Reformulate as (independent of normalizing constant of *q_i*)

$$\max_{\boldsymbol{v}_i \in \mathcal{P}} \mathbb{E}_p[\mathbb{I}_{t_i}^{(1)} \log(q_{\boldsymbol{v}_i})]$$

▶ Solve approximately by replacing \mathbb{E}_p with Monte Carlo estimator

$$\max_{\boldsymbol{v}_i \in \mathcal{P}} \frac{1}{m} \sum_{i=1}^m \mathbb{I}_{t_i}^{(1)}(Z_i) \log(q_{\boldsymbol{v}_i}(Z_i)), \qquad Z_1, \ldots, Z_m \sim p$$

 \Rightarrow Optimization problems affected by rareness of event $\mathbb{I}_{t_i}^{(1)}(Z)$ [Rubinstein, 1999], [Rubinstein, 2001]

MFCE: Iterative cross-entropy procedure

Reuse $q_{v_{i-1}}$ as biasing density for constructing q_{v_i}

$$\max_{\boldsymbol{v}_i \in \mathcal{P}} \mathbb{E}_{q_{\boldsymbol{v}_{i-1}}} \left[\mathbb{I}_{t_i}^{(1)} \frac{p}{q_{\boldsymbol{v}_{i-1}}} \log(q_{\boldsymbol{v}_i}) \right]$$

• Choose
$$t_1 \gg t$$
, solve for $v_1 \in \mathcal{P}$ with

$$\max_{\boldsymbol{v_1} \in \mathcal{P}} \mathbb{E}_{\boldsymbol{\rho}}[\mathbb{I}_{t_1}^{(1)} \log(q_{\boldsymbol{v_1}})]$$

• Select
$$t_2 < t_1$$
, solve for $\boldsymbol{v}_2 \in \mathcal{P}$ with

$$\max_{\boldsymbol{v}_{2} \in \mathcal{P}} \mathbb{E}_{q_{\boldsymbol{v}_{1}}} [\mathbb{I}_{t_{2}}^{(1)} \frac{\boldsymbol{\rho}}{q_{\boldsymbol{v}_{1}}} \log(q_{\boldsymbol{v}_{2}})]$$

• Repeat until threshold t is reached and parameter \mathbf{v}_* is obtained

• Reweighed optimization problems have same optimum as original problems Once q_{v_*} has been found, cross-entropy estimator of P_f is

$$\hat{P}_t^{\mathsf{CE}} = \frac{1}{m} \sum_{i=1}^m \mathbb{I}_t^{(1)}(Z_i^*) \frac{p(Z_i^*)}{q_{\boldsymbol{\nu}_*}(Z_i^*)}, \qquad Z_1^*, \dots, Z_m^* \sim q_{\boldsymbol{\nu}_*}$$

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MFCE: Costs of the cross-entropy method

Step t_i defined by quantile parameter ρ

- Quantile parameter typically $ho \in [10^{-2}, 10^{-1}]$
- Step t_i is ρ quantile corresponding to $q_{v_{i-1}}$

Number of steps T

- Introduce minimal step size $\delta > 0$
- Number steps T is bounded as

$$T \leq rac{t_1 - t}{\delta}$$

Costs of cross-entropy method bounded as

$$c(\hat{P}_t^{\mathsf{CE}}) \leq rac{(t_1 - t) \, mw_1}{\delta}$$



 \Rightarrow Critically depends on t_1 and thus on density used in first step

Our approach: Use surrogates to reduce t_1 and thus #iters with $f^{(1)}$

 $f^{(2)},\ldots,f^{(k)}:\mathcal{Z}\to\mathcal{Y}$



- Find biasing density $q_{\mathbf{v}_*}^{(k)}$ with surrogate $f^{(k)}$
- ▶ Start with $q_{v_*}^{(k)}$ to find biasing density $q_{v_*}^{(k-1)}$ with $f^{(k-1)}$
- Repeat until $q_{\mathbf{v}_*}^{(1)}$ is found with $f^{(1)}$

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MFCE: Analysis for two models

Consider high-fidelity $f^{(1)}$ and low-fidelity model $f^{(2)}$

- Let $q_{\mathbf{v}_*}^{(2)}$ be the biasing density found with $f^{(2)}$
- Set $t_1^{(1)}$ to be the ρ -quantile corresponding to $q_{\mathbf{v}_*}^{(2)}$
- Set t_p to be the ρ -quantile corresponding to the nominal density p
- Need to show

$$\mathbb{P}_{q_{v_{*}}^{(2)}}\left[f^{(1)} \leq t_{1}^{(1)}\right] - \mathbb{P}_{\rho}\left[f^{(1)} \leq t_{1}^{(1)}\right] \geq 0$$

to have $t_1^{(1)} \leq t_p$

Make two assumptions

- A "local" bound $0 < \alpha$ on the error $|f^{(1)}(z) f^{(2)}(z)|$
- Lipschitz continuous distribution functions $F_q^{(i)}(t) = \mathbb{P}_q[f^{(i)} \leq t]$

Proposition 2 in [P., Kramer, Willcox, 2017]

$$\mathbb{P}_{q_{\mathbf{v}_{*}}^{(2)}}\left[f^{(1)} \le t_{1}^{(1)}\right] - \mathbb{P}_{p}\left[f^{(1)} \le t_{1}^{(1)}\right] \gtrsim -\alpha$$

MFCE: Numerical examples: Heat transfer

Consider

$$-\nabla \cdot (a(\omega,\xi)\nabla u(\omega,\xi)) = 1, \qquad \xi \in (0,1), \qquad (1)$$

$$u(\omega,0) = 0, \qquad (2$$

$$\partial_n u(\omega, 1) = 0, \qquad (3)$$

Coefficient a is given as

$$a(\omega,\xi) = e^{z_1(\omega)} e^{-0.5\frac{|\xi-0.5|}{0.0225}} + e^{z_2(\omega)} e^{-0.5\frac{|\xi-0.8|}{0.0225}}$$

- Random vector $Z = [z_1, z_2]$ normally distributed
- System response

$$f(Z(\omega)) = u(\omega, 1)$$

- ▶ Discretize with varying mesh width $h \in \{2^{-8}, 2^{-7}, \dots, 2^{-4}\}$
- Obtain models $f^{(1)}, \ldots, f^{(5)}$

Goal is to estimate $P_f = \mathbb{P}_p[f^{(1)} \le 1.4]$ with reference $P_f \approx 6 \times 10^{-9}$

MFCE: Numerical examples: Speedup



- Single-fidelity approach uses f⁽¹⁾
- ▶ MFCE uses *f*⁽¹⁾,...,*f*⁽⁵⁾
- MFCE reduces runtime by almost 2 orders of magnitude

MFCE: Numerical examples: Number of iterations



- Number of iterations averaged over 30 runs
- MFCE performs most iterations with coarse-grid models

MFCE: Numerical examples: Reacting flow



Reacting-flow problem

- Three models: data-fit, projection-based, and high-fidelity model
- Estimate probability that temperature is below threshold, ref $P_f \approx 10^{-6}$
- MFCE reduces iterations with high-fidelity model from pprox 5 to pprox 1

MFCE: Apply to OpenAeroStruct

Consider baseline UAV definition in OpenAeroStruct

- Design variables are thickness and position of control points
- Uncertain flight conditions (angle of attack, air density, Mach number)
- Output is fuel burn

```
Estimate 10^{-6}-quantile \gamma (value-at-risk)
```

$$\mathbb{P}_{\rho}[f^{(1)} \leq \gamma] = 10^{-6}$$

Derive a data-fit surrogate

- \blacktriangleright Take a 3 \times 3 \times 3 equidistant grid in stochastic domain
- Evaluate high-fidelity model at those 27 points
- Derive linear interpolant of output

Apply multifidelity pre-conditioned cross-entropy method

MFCE: OpenAeroStruct: Value-at-Risk computation



Computing 10⁻⁶-quantile for a fixed design point

Multifidelity approach achieves up to one order of magnitude speedup

Conclusions



Multifidelity rare event simulation

- Leverage surrogate models for runtime speedup
- Recourse to high-fidelity model for accuracy guarantees

Our references

- 1 P., Cui, Marzouk, Willcox: *Multifidelity Importance Sampling*. Computer Methods in Applied Mechanics and Engineering, 300:490-509, 2016.
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PostDoc position available

data-driven model reduction
uncertainty propagation
Bayesian inverse problems

https://pehersto.engr.wisc.edu peherstorfer@wisc.edu